Ensemble discrimination via selective random rotations and projective measurements ¹

C S Sudheer Kumar

NMR Research Center Department of Physics IISER Pune



Young Quantum-2017 HRI, Allahabad

February 27, 2017

◆□▶ ◆□▶ ▲□▶ ▲□▶ □ のQ@

¹viXra:1702.0086



Problem Definition and Introduction

Technical Details

Numerical Simulation Results

Conclusion

◆□▶ ◆□▶ ◆ □▶ ◆ □▶ ● □ ● ● ●

What is the problem?

Ensemble \mathcal{E}_1

$$N_{|0
angle}$$
 number of $|0
angle s$ and $N_{|1
angle}$ number of $|1
angle s$

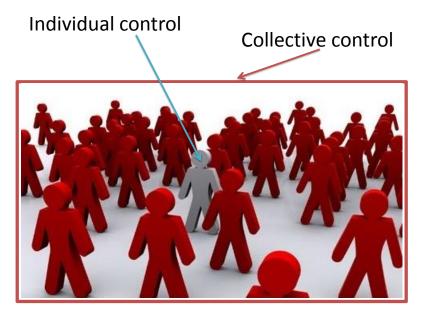
Ensemble \mathcal{E}_{γ} $N_{|+\rangle}$ number of $|+\rangle s$ and $N_{|-\rangle}$ number of $|-\rangle s$ $N_{|+\rangle} + N_{|-\rangle} = N,$ $N_{|+\rangle} / N \cong 1/2, \ N_{|-\rangle} / N \cong 1/2,$

 $|\pm\rangle = (|0\rangle \pm |1\rangle) / \sqrt{2}.$

 $N_{|0\rangle} + N_{|1\rangle} = N,$ $N_{|0\rangle} / N \cong 1/2, \quad N_{|1\rangle} / N \cong 1/2, \qquad N_{|1\rangle}$ $|0\rangle, |1\rangle$ are eigenkets of σ_z .

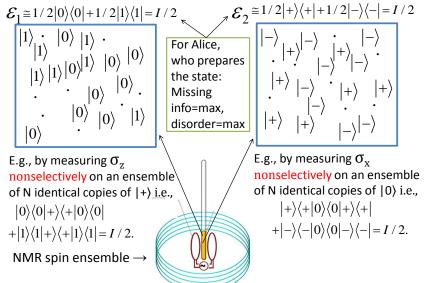
We need to discriminate between ε_1 and ε_2 .

Individual v/s Collective control



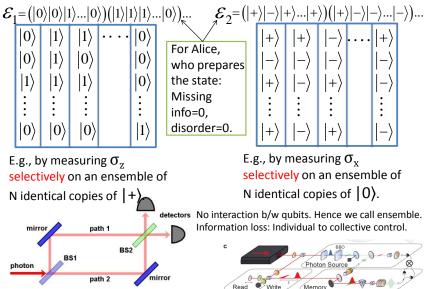
Collective control - Nonselective operations((non)U)

Cannot address and control each of the *N* qubits in the ensemble separately:



Individual control - Selective operations((non)U)

Can address and control each of the *N* qubits in the ensemble separately (\equiv going to 2^{*N*}-D Hilbert space):



C*

$$\begin{array}{l} \bullet \quad \langle \sigma_z \rangle_{|\psi\rangle} = p^+ \times (+1) + p^- \times (-1), \\ (\Delta \sigma_z)^2_{|\psi\rangle} = \langle \sigma_z^2 \rangle_{|\psi\rangle} - \langle \sigma_z \rangle^2_{|\psi\rangle}. \end{array}$$

► $S = (T^+ - T^-)/M$, $\langle S \rangle = \langle \sigma_z \rangle_{|\psi\rangle}$, $\Delta S^2 = (\Delta \sigma_z)^2_{|\psi\rangle}/M$. As *M* increases, ΔS^2 decreases, and hence *S* approaches $\langle \sigma_z \rangle_{|\psi\rangle}$.

- Hence we implicitly neglect variance ΔS^2 .
- We will show that, even though variances got via *E*₁ and *E*₂ tends to zero, their ratio does not tend to one, due to reduction of variance got via *E*₁.
- N_{|0⟩} does not converge to N/2 (but may diverge, we exploit this), where as N_{|0⟩}/N converges to 1/2 as N increases.

$$\begin{array}{l} \flat \ \langle \sigma_z \rangle_{|\psi\rangle} = \rho^+ \times (+1) + \rho^- \times (-1), \\ (\Delta \sigma_z)^2_{|\psi\rangle} = \langle \sigma_z^2 \rangle_{|\psi\rangle} - \langle \sigma_z \rangle^2_{|\psi\rangle}. \end{array}$$

- S = (T⁺ − T⁻)/M, ⟨S⟩ = ⟨σ_z⟩_{|ψ⟩}, ΔS² = (Δσ_z)²_{|ψ⟩}/M. As M increases, ΔS² decreases, and hence S approaches ⟨σ_z⟩_{|ψ⟩}.
- Hence we implicitly neglect variance ΔS^2 .
- We will show that, even though variances got via *E*₁ and *E*₂ tends to zero, their ratio does not tend to one, due to reduction of variance got via *E*₁.
- N_{|0⟩} does not converge to N/2 (but may diverge, we exploit this), where as N_{|0⟩}/N converges to 1/2 as N increases.

•
$$\langle \sigma_z \rangle_{|\psi\rangle} = p^+ \times (+1) + p^- \times (-1),$$

 $(\Delta \sigma_z)^2_{|\psi\rangle} = \langle \sigma_z^2 \rangle_{|\psi\rangle} - \langle \sigma_z \rangle^2_{|\psi\rangle}.$
• $S = (T^+ - T^-)/M, \langle S \rangle = \langle \sigma_z \rangle_{|\psi\rangle}, \Delta S^2 = (\Delta \sigma_z)^2_{|\psi\rangle}/M.$ As
M increases, ΔS^2 decreases, and hence *S* approaches
 $\langle \sigma_z \rangle_{|\psi\rangle}.$

- Hence we implicitly neglect variance ΔS^2 .
- We will show that, even though variances got via E₁ and E₂ tends to zero, their ratio does not tend to one, due to reduction of variance got via E₁.
- N_{|0⟩} does not converge to N/2 (but may diverge, we exploit this), where as N_{|0⟩}/N converges to 1/2 as N increases.

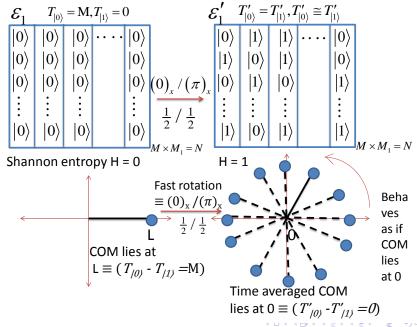
$$\langle \sigma_z \rangle_{|\psi\rangle} = p^+ \times (+1) + p^- \times (-1), (\Delta \sigma_z)^2_{|\psi\rangle} = \langle \sigma_z^2 \rangle_{|\psi\rangle} - \langle \sigma_z \rangle^2_{|\psi\rangle}.$$

- ► $S = (T^+ T^-)/M$, $\langle S \rangle = \langle \sigma_z \rangle_{|\psi\rangle}$, $\Delta S^2 = (\Delta \sigma_z)^2_{|\psi\rangle}/M$. As *M* increases, ΔS^2 decreases, and hence *S* approaches $\langle \sigma_z \rangle_{|\psi\rangle}$.
- Hence we implicitly neglect variance ΔS^2 .
- We will show that, even though variances got via *E*₁ and *E*₂ tends to zero, their ratio does not tend to one, due to reduction of variance got via *E*₁.
- N_{|0⟩} does not converge to N/2 (but may diverge, we exploit this), where as N_{|0⟩}/N converges to 1/2 as N increases.

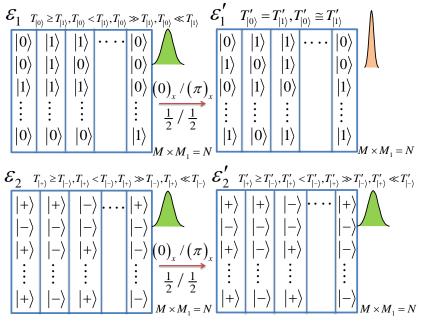
$$\langle \sigma_z \rangle_{|\psi\rangle} = \boldsymbol{p}^+ \times (+1) + \boldsymbol{p}^- \times (-1), (\Delta \sigma_z)^2_{|\psi\rangle} = \langle \sigma_z^2 \rangle_{|\psi\rangle} - \langle \sigma_z \rangle^2_{|\psi\rangle}.$$

- S = (T⁺ − T⁻)/M, ⟨S⟩ = ⟨σ_z⟩_{|ψ⟩}, ΔS² = (Δσ_z)²_{|ψ⟩}/M. As M increases, ΔS² decreases, and hence S approaches ⟨σ_z⟩_{|ψ⟩}.
- Hence we implicitly neglect variance ΔS^2 .
- We will show that, even though variances got via *E*₁ and *E*₂ tends to zero, their ratio does not tend to one, due to reduction of variance got via *E*₁.
- N_{|0⟩} does not converge to N/2 (but may diverge, we exploit this), where as N_{|0⟩}/N converges to 1/2 as N increases.

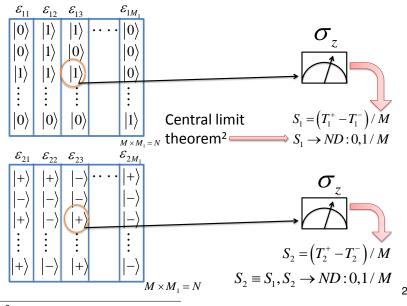
Hypothetical extreme case



Reduction in variance of sample mean (σ_z)

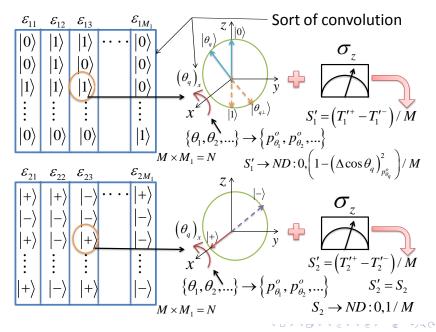


Technical details: Notation and definition



²S. Ross, A first course in probability (Pearson, 2012).

Technical details: Notation and definition



- Note that even though no change in mean upon applying (θ_q)_xs, variance has reduced.
- Measuring σ_z selectively on $|\theta_q\rangle$ s and $|\theta_{q\perp}\rangle$ s is equivalent to tossing differently biased coins (i.e., independently distributed (id) random variables).
- Applying central limit theorem to independently distributed (id) random variables ³, we obtain effective mean

$$\mu_{eff} = \sum_{q} (p_q \langle \sigma_z \rangle_{|\theta_q\rangle} + p_{q\perp} \langle \sigma_z \rangle_{|\theta_{q\perp}\rangle})$$
$$= \sum_{q} (p_q - p_{q\perp}) \cos \theta_q.$$
(1)

where $p_q = M'_q(T_1^+, p_{\theta_q})/M$ ($p_{q\perp} = M'_{q\perp}(T_1^-, p_{\theta_q})/M$), $M'_q(M'_{q\perp})$ is the total number of $|\theta_q\rangle$ s ($|\theta_{q\perp}\rangle$ s). $\sum_q (M'_q + M'_{q\perp}) = M$.

³S. Ross, A first course in probability (Pearson, 201교). 👍 🗸 🗉 🖌 🛓 🔊 ५ 🐑

- Note that even though no change in mean upon applying (θ_q)_xs, variance has reduced.
- ► Measuring σ_z selectively on |θ_q⟩s and |θ_{q⊥}⟩s is equivalent to tossing differently biased coins (i.e., independently distributed (id) random variables).
- Applying central limit theorem to independently distributed (id) random variables ³, we obtain effective mean

$$\mu_{eff} = \sum_{q} (p_q \langle \sigma_z \rangle_{|\theta_q\rangle} + p_{q\perp} \langle \sigma_z \rangle_{|\theta_{q\perp}\rangle})$$
$$= \sum_{q} (p_q - p_{q\perp}) \cos \theta_q. \tag{1}$$

where $p_q = M'_q(T_1^+, p_{\theta_q})/M$ ($p_{q\perp} = M'_{q\perp}(T_1^-, p_{\theta_q})/M$), $M'_q(M'_{q\perp})$ is the total number of $|\theta_q\rangle$ s ($|\theta_{q\perp}\rangle$ s). $\sum_q (M'_q + M'_{q\perp}) = M$.

³S. Ross, A first course in probability (Pearson, 201교). 👍 🗸 🗉 🖌 🛓 🔊 ५ 🐑

- Note that even though no change in mean upon applying (θ_q)_xs, variance has reduced.
- ► Measuring σ_z selectively on |θ_q⟩s and |θ_{q⊥}⟩s is equivalent to tossing differently biased coins (i.e., independently distributed (id) random variables).
- Applying central limit theorem to independently distributed (id) random variables ³, we obtain effective mean

$$\mu_{eff} = \sum_{q} (p_{q} \langle \sigma_{z} \rangle_{|\theta_{q}\rangle} + p_{q\perp} \langle \sigma_{z} \rangle_{|\theta_{q\perp}\rangle})$$
$$= \sum_{q} (p_{q} - p_{q\perp}) \cos \theta_{q}. \tag{1}$$

where $p_q = M'_q(T_1^+, p_{\theta_q})/M$ ($p_{q\perp} = M'_{q\perp}(T_1^-, p_{\theta_q})/M$), $M'_q(M'_{q\perp})$ is the total number of $|\theta_q\rangle s$ ($|\theta_{q\perp}\rangle s$). $\sum_q (M'_q + M'_{q\perp}) = M$.

³S. Ross, A first course in probability (Pearson, 2012). (→ (=) (=) (=) (→ ())

Using Bayes rule

$$p_q = p_1^+ p_{\theta_q}, p_{q\perp} = p_1^- p_{\theta_q}, \qquad (2)$$

where $p_1^{\pm} = T_1^{\pm}/M$, and $p_{\theta_q} = m_q/M$, m_q is the total number of times $(\theta_q)_x$ is applied, $\sum_q m_q = M$.

$$\Rightarrow \mu_{eff} = S_1 \langle \cos \theta_q \rangle_{\rho_{\theta_q}} \tag{3}$$

where $\langle \cos \theta_q \rangle_{p_{\theta_q}} = \sum_q p_{\theta_q} \cos \theta_q$.

►
$$S_1 \rightarrow \text{ND} : 0, 1/M,$$

 $p_1^{\pm} \rightarrow \text{ND} : 1/2, 1/(4M)(:: T_1^{\pm} \rightarrow \text{ND} : M/2, M/4), \text{ and}$
 $p_{\theta_q} \rightarrow \text{ND} : p_{\theta_q}^o, \sigma_{m_q}^2/M^2(:: m_q \rightarrow \text{ND} : p_{\theta_q}^oM, \sigma_{m_q}^2)$ where
 $\sigma_{m_q}^2 \sim M.$ Hence we need to take care of the variance
(however small) present in them.

Using Bayes rule

$$p_q = p_1^+ p_{\theta_q}, p_{q\perp} = p_1^- p_{\theta_q}, \qquad (2)$$

where $p_1^{\pm} = T_1^{\pm}/M$, and $p_{\theta_q} = m_q/M$, m_q is the total number of times $(\theta_q)_x$ is applied, $\sum_q m_q = M$.

$$\Rightarrow \mu_{eff} = S_1 \langle \cos \theta_q \rangle_{\rho_{\theta_q}} \tag{3}$$

where $\langle \cos \theta_q \rangle_{p_{\theta_q}} = \sum_q p_{\theta_q} \cos \theta_q$.

► $S_1 \rightarrow \text{ND} : 0, 1/M,$ $p_1^{\pm} \rightarrow \text{ND} : 1/2, 1/(4M)(:: T_1^{\pm} \rightarrow \text{ND} : M/2, M/4), \text{ and}$ $p_{\theta_q} \rightarrow \text{ND} : p_{\theta_q}^o, \sigma_{m_q}^2/M^2(:: m_q \rightarrow \text{ND} : p_{\theta_q}^oM, \sigma_{m_q}^2)$ where $\sigma_{m_q}^2 \sim M.$ Hence we need to take care of the variance (however small) present in them. Using Bayes rule

$$p_q = p_1^+ p_{\theta_q}, p_{q\perp} = p_1^- p_{\theta_q}, \qquad (2)$$

where $p_1^{\pm} = T_1^{\pm}/M$, and $p_{\theta_q} = m_q/M$, m_q is the total number of times $(\theta_q)_x$ is applied, $\sum_q m_q = M$.

$$\Rightarrow \mu_{eff} = S_1 \langle \cos \theta_q \rangle_{\rho_{\theta_q}} \tag{3}$$

where $\langle \cos heta_q
angle_{m{
ho}_q} = \sum_q m{
ho}_{ heta_q} \cos heta_q.$

► $S_1 \rightarrow \text{ND} : 0, 1/M,$ $p_1^{\pm} \rightarrow \text{ND} : 1/2, 1/(4M)(:: T_1^{\pm} \rightarrow \text{ND} : M/2, M/4), \text{ and}$ $p_{\theta_q} \rightarrow \text{ND} : p_{\theta_q}^o, \sigma_{m_q}^2/M^2(:: m_q \rightarrow \text{ND} : p_{\theta_q}^oM, \sigma_{m_q}^2)$ where $\sigma_{m_q}^2 \sim M.$ Hence we need to take care of the variance (however small) present in them. Applying central limit theorem to id random variables, we obtain effective variance ⁴

$$\begin{aligned} (\Delta \sigma_z)_{eff}^2 &= \sum_q (p_q (\Delta \sigma_z)_{|\theta_q\rangle}^2 + p_{q\perp} (\Delta \sigma_z)_{|\theta_{q\perp}\rangle}^2) \\ &= \sum_q (p_q + p_{q\perp}) \sin^2 \theta_q = 1 - \langle \cos^2 \theta_q \rangle_{p_{\theta_q}}, (4) \end{aligned}$$

where
$$(\Delta \sigma_z)^2_{|\theta_q\rangle} = \langle \sigma_z^2 \rangle_{|\theta_q\rangle} - \langle \sigma_z \rangle^2_{|\theta_q\rangle}$$
,
 $\langle \cos^2 \theta_q \rangle_{p_{\theta_q}} = \sum_q p_{\theta_q} \cos^2 \theta_q$.

• $(\Delta \sigma_z)_{eff}^2 \neq \langle \sigma_z^2 \rangle_{\rho'_{1j}} - \langle \sigma_z \rangle_{\rho'_{1j}}^2$ where $\rho'_{1j} = \sum_q (p_q |\theta_q \rangle \langle \theta_q | + p_{q\perp} |\theta_{q\perp} \rangle \langle \theta_{q\perp} |)$, because in going from \mathcal{E}'_{1i} to ρ'_{1j} , there is information loss.

• According to central limit theorem, in the large *M* limit, probability distribution of effective sample mean S'_1 , for given values of p_{θ_q} s and S_1 (i.e., for given values of m_q s and T_1^+ i.e., for a given \mathcal{E}'_{1j}), tends to normal distribution⁴ i.e, $S'_1 \rightarrow \text{ND}$: μ_{eff} , $(\Delta \sigma_z)^2_{eff}/M$.

⁴S. Ross, A first course in probability (Pearson, 2012).

 Applying central limit theorem to id random variables, we obtain effective variance ⁴

$$\begin{aligned} (\Delta \sigma_z)_{eff}^2 &= \sum_q (p_q (\Delta \sigma_z)_{|\theta_q\rangle}^2 + p_{q\perp} (\Delta \sigma_z)_{|\theta_{q\perp}\rangle}^2) \\ &= \sum_q (p_q + p_{q\perp}) \sin^2 \theta_q = 1 - \langle \cos^2 \theta_q \rangle_{p_{\theta_q}}, (4) \end{aligned}$$

where
$$(\Delta \sigma_z)^2_{|\theta_q\rangle} = \langle \sigma_z^2 \rangle_{|\theta_q\rangle} - \langle \sigma_z \rangle^2_{|\theta_q\rangle}$$
,
 $\langle \cos^2 \theta_q \rangle_{\rho_{\theta_q}} = \sum_q p_{\theta_q} \cos^2 \theta_q$.
• $(\Delta \sigma_z)^2_{eff} \neq \langle \sigma_z^2 \rangle_{\rho'_{1j}} - \langle \sigma_z \rangle^2_{\rho'_{1j}}$ where
 $\rho'_{1j} = \sum_q (p_q |\theta_q\rangle \langle \theta_q | + p_{q\perp} |\theta_{q\perp}\rangle \langle \theta_{q\perp} |)$, because in going
from \mathcal{E}'_{1j} to ρ'_{1j} , there is information loss.

• According to central limit theorem, in the large *M* limit, probability distribution of effective sample mean S'_1 , for given values of p_{θ_q} s and S_1 (i.e., for given values of m_q s and T_1^+ i.e., for a given \mathcal{E}'_{1j}), tends to normal distribution⁴ i.e, $S'_1 \rightarrow \text{ND} : \mu_{eff}, (\Delta \sigma_z)^2_{eff}/M$.

⁴S. Ross, A first course in probability (Pearson, 2012).

 Applying central limit theorem to id random variables, we obtain effective variance ⁴

$$\begin{aligned} (\Delta \sigma_z)_{eff}^2 &= \sum_q (p_q (\Delta \sigma_z)_{|\theta_q\rangle}^2 + p_{q\perp} (\Delta \sigma_z)_{|\theta_{q\perp}\rangle}^2) \\ &= \sum_q (p_q + p_{q\perp}) \sin^2 \theta_q = 1 - \langle \cos^2 \theta_q \rangle_{p_{\theta_q}}, (4) \end{aligned}$$

where
$$(\Delta \sigma_z)^2_{|\theta_q\rangle} = \langle \sigma_z^2 \rangle_{|\theta_q\rangle} - \langle \sigma_z \rangle^2_{|\theta_q\rangle}$$
,
 $\langle \cos^2 \theta_q \rangle_{p_{\theta_q}} = \sum_q p_{\theta_q} \cos^2 \theta_q$.
• $(\Delta \sigma_z)^2_{eff} \neq \langle \sigma_z^2 \rangle_{\rho'_{1j}} - \langle \sigma_z \rangle^2_{\rho'_{1j}}$ where
 $\rho'_{1j} = \sum_q (p_q |\theta_q\rangle \langle \theta_q | + p_{q\perp} |\theta_{q\perp}\rangle \langle \theta_{q\perp} |)$, because in going
from \mathcal{E}'_{1j} to ρ'_{1j} , there is information loss.

• According to central limit theorem, in the large *M* limit, probability distribution of effective sample mean S'_1 , for given values of p_{θ_q} s and S_1 (i.e., for given values of m_q s and T_1^+ i.e., for a given \mathcal{E}'_{1j}), tends to normal distribution⁴ i.e., $S'_1 \rightarrow \text{ND} : \mu_{eff}, (\Delta \sigma_z)^2_{eff}/M$.

⁴S. Ross, A first course in probability (Pearson, 2012). (♂) (3) (3) (3)

Resultant probability density of S'₁ is given by

$$f(S'_{1}) = \int \prod_{i,i\neq l} \{ dp_{\theta_{i}}(\mathsf{Nd}(p_{\theta_{i}}) : p^{o}_{\theta_{i}}, \sigma^{2}_{m_{i}}/M^{2}) \}$$

$$\times \left(\operatorname{Nd}(S_1') : S_1 \langle \cos \theta_q \rangle_{p_{\theta_q}}, (1 - \langle \cos^2 \theta_q \rangle_{p_{\theta_q}}) / M \right), \quad (5)$$

where $(Nd(x) : \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma}} \exp(-(x - \mu)^2/(2\sigma^2))$ (i.e., Normal probability density function with mean μ and variance σ^2), $dx(Nd(x) : \mu, \sigma^2)$ is the probability of obtaining value *x* of normally distributed random variable *x*. • Integrating out S_1 we get

$$f(S'_{1}) = \int \prod_{i,i\neq l} \{ dp_{\theta_{i}}(\operatorname{Nd}(p_{\theta_{i}}) : p_{\theta_{i}}^{o}, \sigma_{m_{i}}^{2}/M^{2}) \}$$
$$\times (\operatorname{Nd}(S'_{1}) : 0, (1 - (\Delta \cos \theta_{q})_{p_{\theta_{q}}}^{2})/M), \qquad (6)$$
ere $(\Delta \cos \theta_{q})_{p_{\theta_{q}}}^{2} = \langle \cos^{2} \theta_{q} \rangle_{p_{\theta_{q}}} - \langle \cos \theta_{q} \rangle_{p_{\theta_{q}}}^{2}.$

Resultant probability density of S'₁ is given by

wł

$$f(S'_{1}) = \int \prod_{i,i \neq l} \{ dp_{\theta_{i}}(\mathsf{Nd}(p_{\theta_{i}}) : p_{\theta_{i}}^{o}, \sigma_{m_{i}}^{2}/M^{2}) \}$$

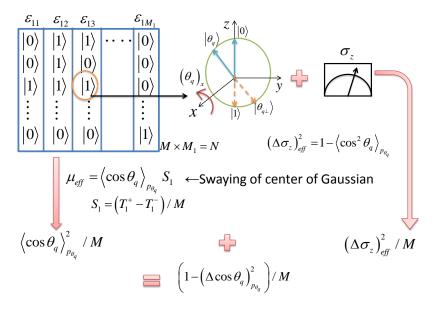
$$\times dS_1(\operatorname{Nd}(S_1):0,1/M) \times \left(\operatorname{Nd}(S_1'):S_1\langle \cos\theta_q\rangle_{p_{\theta_q}},(1-\langle \cos^2\theta_q\rangle_{p_{\theta_q}})/M\right), \quad (5)$$

where $(Nd(x) : \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma}} \exp(-(x - \mu)^2/(2\sigma^2))$ (i.e., Normal probability density function with mean μ and variance σ^2), $dx(Nd(x) : \mu, \sigma^2)$ is the probability of obtaining value *x* of normally distributed random variable *x*. Integrating out *S*₁ we get

$$f(S'_{1}) = \int \prod_{i,i\neq l} \{ dp_{\theta_{i}}(\operatorname{Nd}(p_{\theta_{i}}) : p_{\theta_{i}}^{o}, \sigma_{m_{i}}^{2}/M^{2}) \} \\ \times (\operatorname{Nd}(S'_{1}) : 0, (1 - (\Delta \cos \theta_{q})_{p_{\theta_{q}}}^{2})/M), \qquad (6)$$

here $(\Delta \cos \theta_{q})_{p_{\theta_{q}}}^{2} = \langle \cos^{2} \theta_{q} \rangle_{p_{\theta_{q}}} - \langle \cos \theta_{q} \rangle_{p_{\theta_{q}}}^{2}.$

Resultant variance



Consider θ_q = θ₀, ∀q (i.e., no randomness). Then Eq. (6) reduces to f(S'₁) = (Nd(S'₁) : 0, (1 − 0)/M) = g(S'₂), hence no discrimination.

Consider the simplest case: {θ₁, θ₂} → {p^o_{θ1}, p^o_{θ2}}.

$$\Rightarrow f(S_1') \approx \int_{p_{\theta_1}^o - \epsilon}^{p_{\theta_1}^o + \epsilon} dp_{\theta_1} \delta(p_{\theta_1} - p_{\theta_1}^o)$$
$$\times (\operatorname{Nd}(S_1') : 0, (1 - p_{\theta_1}(1 - p_{\theta_1})(\cos \theta_1 - \cos \theta_2)^2)/M)$$
$$= (\operatorname{Nd}(S_1') : 0, (1 - (\Delta \cos \theta_q)_{p_{\theta_q}^o}^2)/M), \quad (7)$$

where $\epsilon > 0$ (\cdot : no swaying of center of Gaussian in Eq. (6)).

- Consider θ_q = θ₀, ∀q (i.e., no randomness). Then Eq. (6) reduces to f(S'₁) = (Nd(S'₁) : 0, (1 − 0)/M) = g(S'₂), hence no discrimination.
- Consider the simplest case: $\{\theta_1, \theta_2\} \rightarrow \{p_{\theta_1}^o, p_{\theta_2}^o\}$.

$$\Rightarrow f(S'_1) \approx \int_{p_{\theta_1}^o - \epsilon}^{p_{\theta_1}^o + \epsilon} dp_{\theta_1} \delta(p_{\theta_1} - p_{\theta_1}^o)$$
$$\times (\operatorname{Nd}(S'_1) : 0, (1 - p_{\theta_1}(1 - p_{\theta_1})(\cos \theta_1 - \cos \theta_2)^2)/M)$$
$$= (\operatorname{Nd}(S'_1) : 0, (1 - (\Delta \cos \theta_q)_{p_{\theta_q}^o}^2)/M), \quad (7)$$

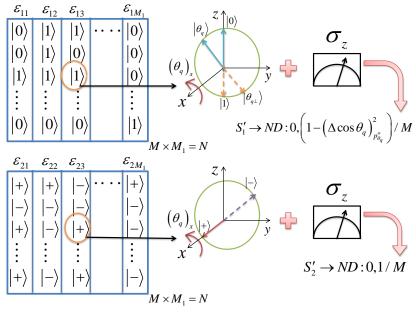
where $\epsilon > 0$ (\cdot : no swaying of center of Gaussian in Eq. (6)).

- Consider θ_q = θ₀, ∀q (i.e., no randomness). Then Eq. (6) reduces to f(S'₁) = (Nd(S'₁) : 0, (1 − 0)/M) = g(S'₂), hence no discrimination.
- Consider the simplest case: $\{\theta_1, \theta_2\} \rightarrow \{p^o_{\theta_1}, p^o_{\theta_2}\}.$

$$\Rightarrow f(S_1') \approx \int_{p_{\theta_1}^o - \epsilon}^{p_{\theta_1}^o + \epsilon} dp_{\theta_1} \delta(p_{\theta_1} - p_{\theta_1}^o)$$
$$\times (\operatorname{Nd}(S_1') : 0, (1 - p_{\theta_1}(1 - p_{\theta_1})(\cos \theta_1 - \cos \theta_2)^2)/M)$$
$$= (\operatorname{Nd}(S_1') : 0, (1 - (\Delta \cos \theta_q)_{p_{\theta_q}^o}^2)/M), \quad (7)$$

where $\epsilon > 0$ (\cdot : no swaying of center of Gaussian in Eq. (6)).

Summary so for



지나 제지막 제품 제품 제품 제품 가지()*

Nonlinearity in action

We will show how nonlinearity is reducing the variance. We have ∆S'²₁ ≈ (1 - (∆ cos θ_q)²_{p^o_{θq}})/M = (⟨cos θ_q⟩²_{p^o_{θq}} + ⟨sin² θ_q⟩_{p^o_{θq}})/M (Eq. (7)).
Let {θ₁(= 0), θ₂(= π/2)} → {p^o₀, p^o_{π/2}}.
⇒ ∆S'²₁ ≈ [(p^o₀ cos 0 + p^o_{π/2} cos(π/2))² + p^o₀ sin² 0 + p^o_{π/2} sin²(π/2)]/M = (p^{o2}₀ + p^o_{π/2})/M < 1/M.

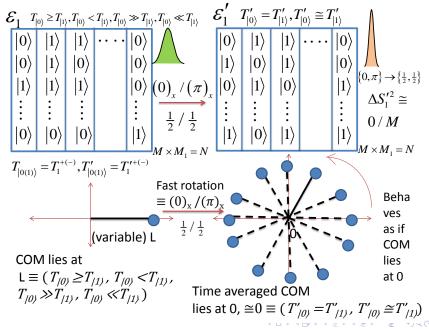
Nonlinearity in action

We will show how nonlinearity is reducing the variance. We have ∆S'₁² ≈ (1 - (∆ cos θ_q)²_{p⁰_{θq}})/M = (⟨cos θ_q⟩²_{p⁰_{θq}} + ⟨sin² θ_q⟩_{p⁰_{θq}})/M (Eq. (7)).
Let {θ₁(= 0), θ₂(= π/2)} → {p⁰₀, p⁰_{π/2}}.
⇒ ∆S'₁² ≈ [(p⁰₀ cos 0 + p⁰_{π/2} cos(π/2))² + p⁰₀ sin² 0 + p⁰_{0/2} sin²(π/2)]/M = (p⁰₀² + p⁰₋₍₂)/M < 1/M.

Nonlinearity in action

We will show how nonlinearity is reducing the variance. We have ∆S'²₁ ≈ (1 - (∆ cos θ_q)²_{p^o_{θq}})/M = (⟨cos θ_q⟩²_{p^o_{θq}} + ⟨sin² θ_q⟩_{p^o_{θq}})/M (Eq. (7)).
Let {θ₁(= 0), θ₂(= π/2)} → {p^o_{θq}, p^o_{π/2}}.
⇒ ∆S'²₁ ≈ [(p^o₀ cos 0 + p^o_{π/2} cos(π/2))² + p^o₀ sin² 0 + p^o_{π/2} sin²(π/2)]/M = (p^{o2}₀ + p^o_{π/2})/M < 1/M.

Smoothing out nonuniformities



MATLAB simulation results

- MATLAB generates standard uniformly distributed Pseudo Random Numbers (PRN) drawn from the open interval (0, 1).
- We want to simulate σ_z measurement on $|\chi\rangle = \cos(\theta/2)|0\rangle + e^{i\phi}\sin(\theta/2)|1\rangle$.
- If we get a PRN in the interval (0, cos²(θ/2)), then it is equivalent to getting outcome +1. Else -1.

(日) (日) (日) (日) (日) (日) (日)

• We simulated the case $\{\theta_1(=0), \theta_2(=\pi)\} \rightarrow \{p^o_{\theta_1}(=1/2), p^o_{\theta_2}(=1/2)\}$

 $\blacktriangleright \Rightarrow \Delta S_1^{\prime 2} \cong 0/M.$

- MATLAB generates standard uniformly distributed Pseudo Random Numbers (PRN) drawn from the open interval (0,1).
- We want to simulate σ_z measurement on $|\chi\rangle = \cos(\theta/2)|0\rangle + e^{i\phi}\sin(\theta/2)|1\rangle$.
- If we get a PRN in the interval (0, cos²(θ/2)), then it is equivalent to getting outcome +1. Else -1.

(日) (日) (日) (日) (日) (日) (日)

• We simulated the case $\{\theta_1(=0), \theta_2(=\pi)\} \rightarrow \{p^o_{\theta_1}(=1/2), p^o_{\theta_2}(=1/2)\}$

 $\blacktriangleright \Rightarrow \Delta S_1^{\prime 2} \cong 0/M.$

- MATLAB generates standard uniformly distributed Pseudo Random Numbers (PRN) drawn from the open interval (0, 1).
- We want to simulate σ_z measurement on $|\chi\rangle = \cos(\theta/2)|0\rangle + e^{i\phi}\sin(\theta/2)|1\rangle.$
- If we get a PRN in the interval (0, cos²(θ/2)), then it is equivalent to getting outcome +1. Else −1.

• We simulated the case $\{\theta_1(=0), \theta_2(=\pi)\} \rightarrow \{p^o_{\theta_1}(=1/2), p^o_{\theta_2}(=1/2)\}.$

 $\blacktriangleright \Rightarrow \Delta S_1^{\prime 2} \cong 0/M.$

- MATLAB generates standard uniformly distributed Pseudo Random Numbers (PRN) drawn from the open interval (0, 1).
- We want to simulate σ_z measurement on $|\chi\rangle = \cos(\theta/2)|0\rangle + e^{i\phi}\sin(\theta/2)|1\rangle.$
- If we get a PRN in the interval (0, cos²(θ/2)), then it is equivalent to getting outcome +1. Else −1.

(日) (日) (日) (日) (日) (日) (日)

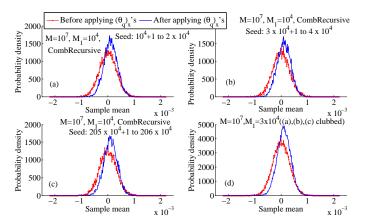
► We simulated the case $\{\theta_1(=0), \theta_2(=\pi)\} \rightarrow \{p_{\theta_1}^o(=1/2), p_{\theta_2}^o(=1/2)\}.$ ► $\Rightarrow \Delta S_1^{\prime 2} \approx 0/M.$

- MATLAB generates standard uniformly distributed Pseudo Random Numbers (PRN) drawn from the open interval (0,1).
- We want to simulate σ_z measurement on $|\chi\rangle = \cos(\theta/2)|0\rangle + e^{i\phi}\sin(\theta/2)|1\rangle$.
- If we get a PRN in the interval (0, cos²(θ/2)), then it is equivalent to getting outcome +1. Else −1.

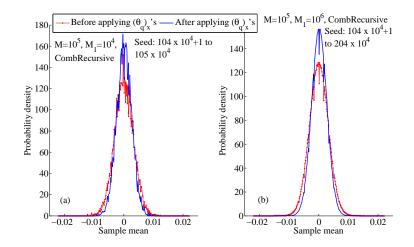
(日) (日) (日) (日) (日) (日) (日)

• We simulated the case $\{\theta_1(=0), \theta_2(=\pi)\} \rightarrow \{p^o_{\theta_1}(=1/2), p^o_{\theta_2}(=1/2)\}.$

 $\blacktriangleright \Rightarrow \Delta S_1^{\prime 2} \cong 0/M.$

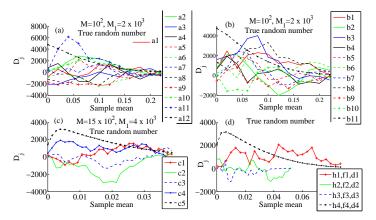


(a) $A_g = 0.6795$ (theory 0.6826895). $A_f = 0.7445$ (theory 1). $A'_g = 0.6795$, and $A'_f = 0.785$. (b) $A_g = 0.6787$, $A_f = 0.739$, $A'_g = 0.6793$, $A'_f = 0.777$. (c) $A_g = 0.685$, $A_f = 0.7492$, $A'_g = 0.6828$, $A'_f = 0.7855$. (d) $A_g = 0.6811$, $A_f = 0.7442$, $A'_g = 0.6811$, $A'_f = 0.7824$.

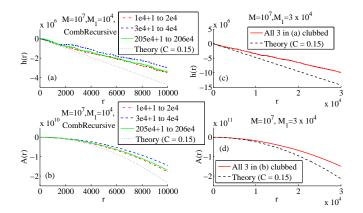


(a) $A_g = 0.6854, A_f = 0.7833$. (b) $A_g = 0.683967, A_f = 0.77997, A'_g = 0.683286, A'_f = 0.780642$.

MATLAB simulation results-True random numbers



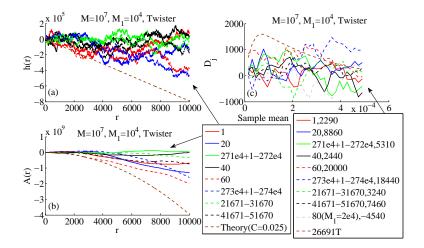
 D_j is the difference in area under the Gaussians ×10⁵ i.e., $D_j = (\sum_{S'_1=-a_j}^{a_j} f(S'_1) - \sum_{S_1=-a_j}^{a_j} g(S_1))\delta S \times 10^5$ where δS is the smallest element (step size) on x-axis (sample mean) considered for plotting, and $a_j = j \times \delta S, j = 1, 2, ...$ Theoretical curve: $(D_i)_{theery}/10, \Delta S'_1^2 \rightarrow 0.1^2/M$ instead of 0/M.



$$h(r) = \sum_{i=1}^{r} (|T_{1i}^{\prime +} - T_{1i}^{\prime -}| - |T_{1i}^{+} - T_{1i}^{-}|), \qquad (8)$$

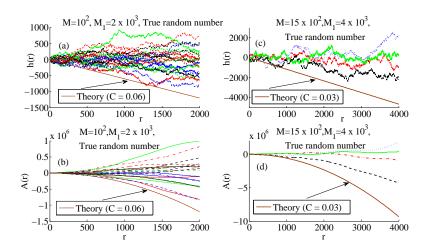
where $r = 1, 2, ..., M_1$.

▲□▶ ▲□▶ ▲□▶ ▲□▶ = 三 のへで



 $26691T: (D_j)_{thoery}/100, \Delta S'^2_1 \rightarrow 0.1^2/M$ instead of 0/M.

MATLAB simulation results-True random numbers



▲□▶▲圖▶▲≣▶▲≣▶ ≣ の�?

- We showed that, if we have individual control, then we can discriminate between two ensembles via selective random rotations, which otherwise (i.e., without individual control) cannot be discriminated, as both are maximally mixed.
- Numerical simulation results support theoretical predictions.
- However the origin of nonlinear effect (reduction in variance) which leads to discrimination is not clear.
- It is interesting to explore whether it is genuine nonlinear effect perhaps due to projective measurement or it is just a consequence of statistical data analysis technique.

- We showed that, if we have individual control, then we can discriminate between two ensembles via selective random rotations, which otherwise (i.e., without individual control) cannot be discriminated, as both are maximally mixed.
- Numerical simulation results support theoretical predictions.
- However the origin of nonlinear effect (reduction in variance) which leads to discrimination is not clear.
- It is interesting to explore whether it is genuine nonlinear effect perhaps due to projective measurement or it is just a consequence of statistical data analysis technique.

(ロ) (同) (三) (三) (三) (○) (○)

- We showed that, if we have individual control, then we can discriminate between two ensembles via selective random rotations, which otherwise (i.e., without individual control) cannot be discriminated, as both are maximally mixed.
- Numerical simulation results support theoretical predictions.
- However the origin of nonlinear effect (reduction in variance) which leads to discrimination is not clear.
- It is interesting to explore whether it is genuine nonlinear effect perhaps due to projective measurement or it is just a consequence of statistical data analysis technique.

- We showed that, if we have individual control, then we can discriminate between two ensembles via selective random rotations, which otherwise (i.e., without individual control) cannot be discriminated, as both are maximally mixed.
- Numerical simulation results support theoretical predictions.
- However the origin of nonlinear effect (reduction in variance) which leads to discrimination is not clear.
- It is interesting to explore whether it is genuine nonlinear effect perhaps due to projective measurement or it is just a consequence of statistical data analysis technique.

- Prof. T S Mahesh: Project initiation and ensemble formulation.
- Prof. R Srikanth: Density matrix formulation.
- Prof. A K Rajagopal: Relative entropy.
- Prof. L Vaidman, Aravinda, Govind Unnikrishnan, Prof. M Ozawa, Deepak Khurana, Anjusha V S.

< □ > < 同 > < 三 > < 三 > < 三 > < ○ < ○ </p>

- Prof. T S Mahesh: Project initiation and ensemble formulation.
- Prof. R Srikanth: Density matrix formulation.
- Prof. A K Rajagopal: Relative entropy.
- Prof. L Vaidman, Aravinda, Govind Unnikrishnan, Prof. M Ozawa, Deepak Khurana, Anjusha V S.

◆□▶ ◆□▶ ▲□▶ ▲□▶ ■ ののの

- Prof. T S Mahesh: Project initiation and ensemble formulation.
- Prof. R Srikanth: Density matrix formulation.
- Prof. A K Rajagopal: Relative entropy.
- Prof. L Vaidman, Aravinda, Govind Unnikrishnan, Prof. M Ozawa, Deepak Khurana, Anjusha V S.

◆□▶ ◆□▶ ▲□▶ ▲□▶ □ のQ@

- Prof. T S Mahesh: Project initiation and ensemble formulation.
- Prof. R Srikanth: Density matrix formulation.
- Prof. A K Rajagopal: Relative entropy.
- Prof. L Vaidman, Aravinda, Govind Unnikrishnan, Prof. M Ozawa, Deepak Khurana, Anjusha V S.

< □ > < 同 > < 三 > < 三 > < 三 > < ○ < ○ </p>

THANK YOU

Single copy picture in 2^N-D Hilbert space

There are two sets: $\mathcal{F}_1 = \{|0\rangle^{\otimes N}, |0\rangle^{\otimes N-1}|1\rangle, ..., |1\rangle^{\otimes N}\}$, and $\mathcal{F}_{2} = \{ |+\rangle^{\otimes N}, |+\rangle^{\otimes N-1} |-\rangle, ..., |-\rangle^{\otimes N} \}.$ \mathcal{F}_i is a complete set of orthonormal basis states in 2^N dimensional (2^N-D) Hilbert space, i = 1, 2. E.g., for N = 2, $\mathcal{F}_1 = \{|0\rangle|0\rangle, |0\rangle|1\rangle, |1\rangle|0\rangle, |1\rangle|1\rangle\}, \text{ and }$ $\mathcal{F}_{2} = \{ |+\rangle |+\rangle, |+\rangle |-\rangle, |-\rangle |+\rangle, |-\rangle |-\rangle \}.$ Let $|\phi_{ii}\rangle \in \mathcal{F}_i$, $i = 1, 2, j = 1, 2, ..., 2^N$. Even though $|\langle \phi_{1i} | \phi_{2k} \rangle|$ tends to zero in the limit $N \to \infty$, $|\phi_{1i}\rangle$ can never become perfectly orthogonal to $|\phi_{2k}\rangle$, because the set \mathcal{F}_i is already complete, i = 1, 2. Hence \mathcal{F}_1 and \mathcal{F}_2 together constitute a set of nontrivial nonorthogonal states. Alice gives Bob, a *single copy* of $|\phi_{ii}\rangle$ chosen with probability $1/2^N$ (i.e., all the states are equally likely to be chosen) from \mathcal{F}_i ,

i = 1 or 2.

Hence $|\phi_{1j}\rangle$ ($|\phi_{2j}\rangle$) is nothing but the renormalized post measurement state of measuring σ_z (σ_x) selectively (i.e., locally) on each of the *N* qubits in the state $|+\rangle^{\otimes N}$ ($|0\rangle^{\otimes N}$), $|0\rangle = (|+\rangle + |-\rangle)/\sqrt{2}$.

Alice tells Bob the way she chose the state from one of $\mathcal{F}_1, \mathcal{F}_2$, but she do not tell him exactly from which set she chose the state. Hence Bob is aware of $\mathcal{F}_1, \mathcal{F}_2$, and Alice's state choosing procedure.

Bob has a single copy of the unknown state $|\phi_{ij}\rangle$, i = 1 or 2. We are going to show that, in the limit $N \to \infty$, even though Bob cannot know the unknown state exactly, still he can know deterministically whether it was chosen from \mathcal{F}_1 or \mathcal{F}_2 (and hence it is deterministic but inexact nonorthogonal state discrimination).

In density matrix formulation, Bob's unknown state is given by:

$$\rho_{i} = \sum_{j=1}^{2^{N}} \frac{1}{2^{N}} |\phi_{ij}\rangle \langle \phi_{ij}| = \frac{\mathbb{1}_{2^{N}}}{2^{N}}, i = 1 \text{ or } 2$$
(9)

where $\mathbb{1}_n$ is $n \times n$ identity matrix. Note that ρ_i represents the state of a single copy of one of $|\phi_{ij}\rangle$ s, $j = 1, 2, ..., 2^N$, which Bob has got, taking into consideration the probability $(1/2^N)$ with which he obtains it. ρ_i represents the state of an ensemble with individual control, but not collective control. Mixedness of ρ_i represents Bob's ignorance about the single copy of the state he has got. Hence it can be purified by selective projective measurement unlike in nonselective ensemble measurement