

# One bin test (on-off problem)

A counting expt. observed events  $\rightarrow n$ ,  $n = \underline{\underline{\mu_s}} + \underline{\underline{\mu_b}}$

$$\underline{E}(n) = \boxed{\mu_s} + \boxed{b}$$

↑ signal strength parameter (p.o.i)

b  $\Rightarrow$  nuisance parameter

s  $\Rightarrow$  comes from theory.

Two hypotheses :  $\mu = 0$  background only  
 $\mu \neq 0$  signal + background.

null

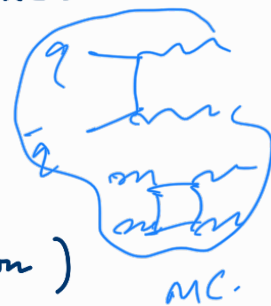
$$\mu = 0 : P(n | b) = \frac{e^{-b} b^n}{n!}$$

$$\mu \neq 0 : P(n | \mu_s + b) = \frac{e^{-\mu_s + b} (\mu_s + b)^n}{n!}$$

b is constrained by a control measurement  
 (let's say, an MC estimate of b)

m counts in control sample.

$$\underline{E}(m) = \underline{\tau} b \quad (\text{Assume, } \underline{\tau} \text{ is known})$$



Joint prob of observing  $(n, m)$

$$L(\mu, b) = \left[ \frac{(\mu_s + b)^n e^{-(\mu_s + b)}}{n!} \right] \left[ \frac{(\tau b)^m e^{-\tau b}}{m!} \right]$$

real run                      auxiliary

When  $\mu \neq 0$

$$\left. \begin{aligned} \frac{\partial L}{\partial \mu} = 0 \\ \frac{\partial L}{\partial b} = 0 \end{aligned} \right\} \Rightarrow$$

$$\hat{\mu} = \frac{n - \frac{m}{\tau}}{s}$$

$$\hat{b} = \frac{m}{\tau}$$

↑ post-fit. estimates

When  $\mu$  is not allowed to vary, but kept at some specified value (profiling)

$$\hat{b} = \frac{(n+m) - (1+\tau)\mu S}{2(1+\tau)} + \left[ \frac{(n+m - (1+\tau)\mu S)^2 + 4(1+\tau)m\mu S}{4(1+\tau)^2} \right]^{1/2}$$

in particular, for  $\mu = 0$

$$\hat{b} = \frac{(n+m)}{2(1+\tau)} + \frac{(n+m)}{2(1+\tau)} = \frac{n+m}{1+\tau}$$

: Weighted average of the two counts.

We construct a test statistic

\* "statistic means a function of data"

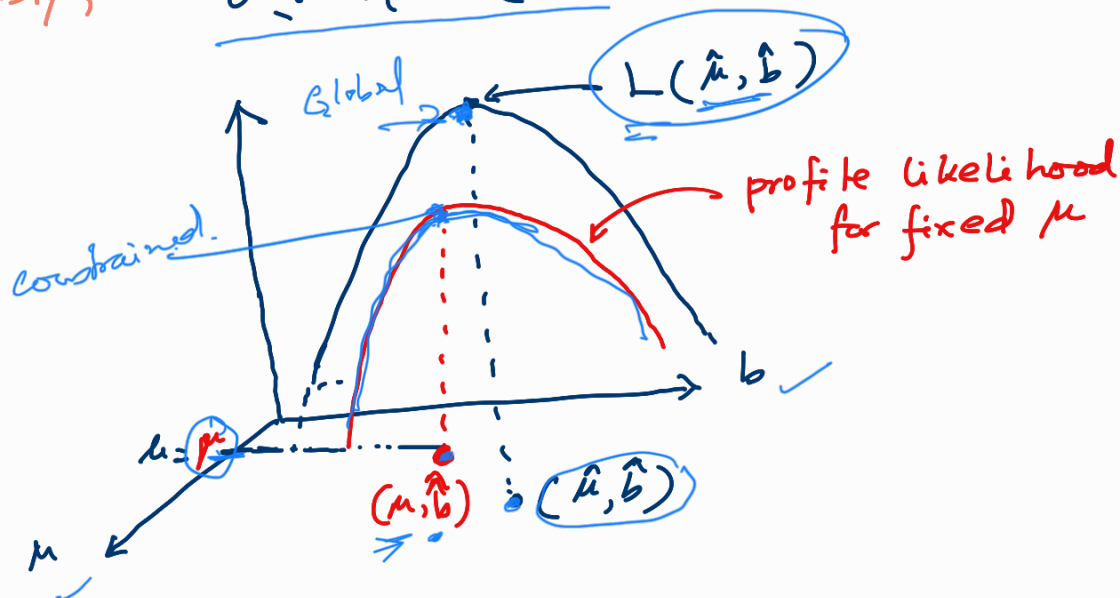
$$\lambda(\mu) = \frac{L(\mu, \hat{b})}{L(\hat{\mu}, \hat{b})}$$

$\mu$

← conditional max  
← unconditional max

profile L.R.

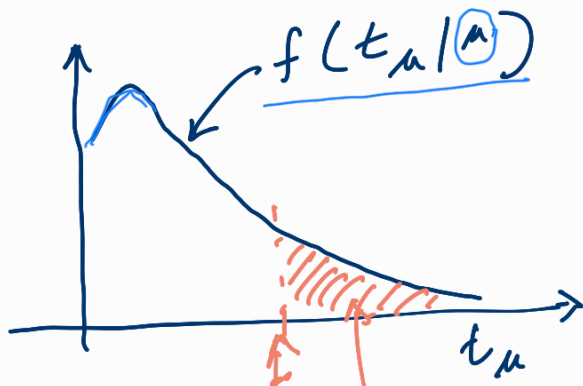
Obviously,  $0 \leq \lambda(\mu) \leq 1$



$$t_\mu = -2 \ln \lambda(\mu)$$

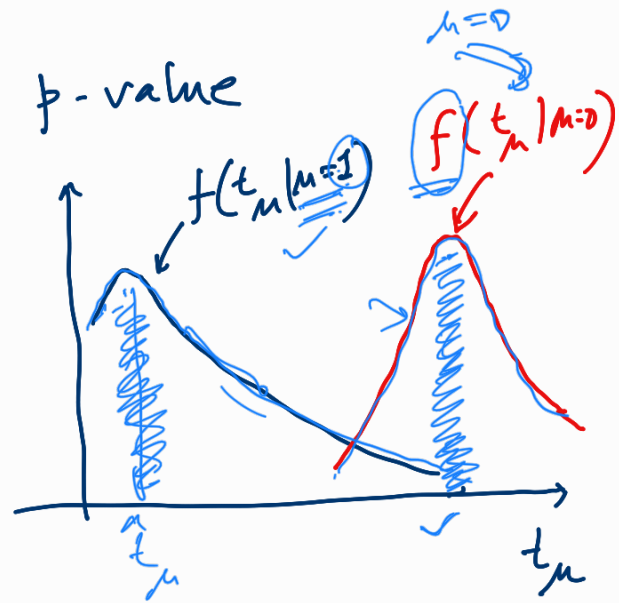
↑ statistic, random.

# Distribution of $t_\mu$ and p-value



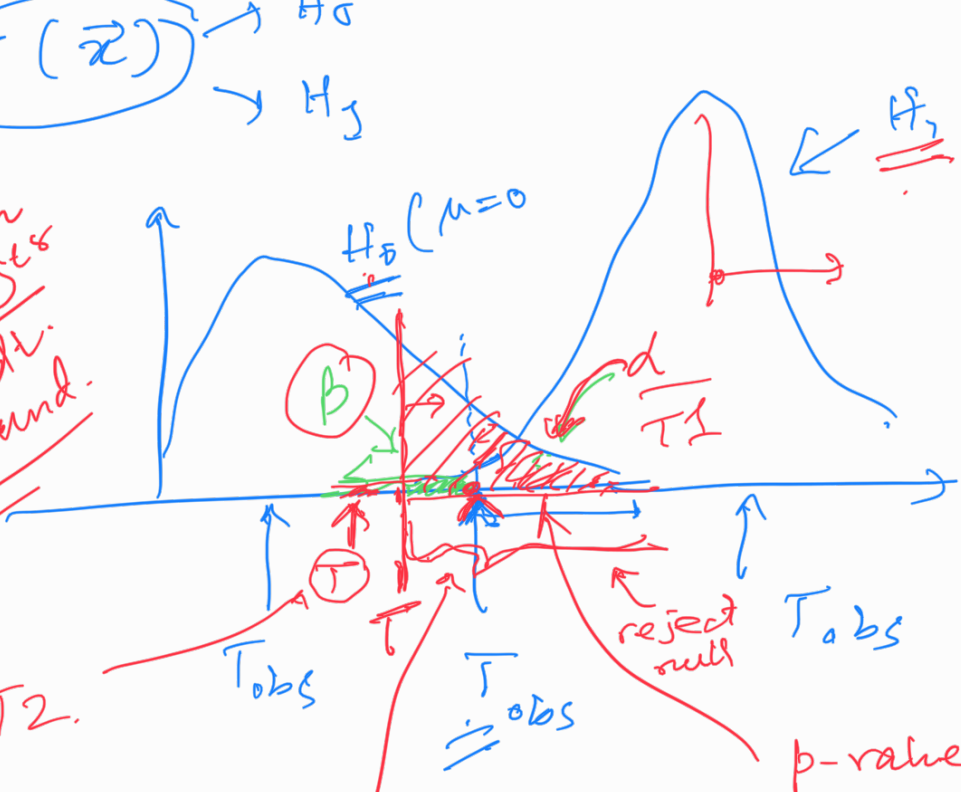
$-0.5$

$$p_\mu = \int_{t_\mu}^{\infty} f(t_\mu | \mu) dt_\mu$$



$T(\bar{x})$   
 $\rightarrow H_0$   
 $\rightarrow H_1$

Casella-Berger  
Brandt-Sigmund



$\alpha$  reject  
 $(1-\beta)$   
largest

Neyman-Pearson  
UMP

critical region.

p-value ( $T_{obs} | H_0$ )  
p-value ( $T_{obs} | H_1$ )